## Advanced Mathematical Methods WS 2021/22

### **5** Hypothesis Testing

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## Readings

 A. Papoulis and A. U. Pillai. Probability, Random Variables and Stochastic Processes.
 Mc Graw Hill, fourth edition, 2002, Chapter 8

### **Online References**

MIT Course on Probabilistic Systems Analysis and Applied Probability (by John Tsitsiklis)

• Lecture 25: Classical Inference III

# Hypothesis testing

#### Testing procedure

step 1: set up a null hypothesis H<sub>0</sub> and an alternative hypothesis H<sub>1</sub>
step 2: construct a test statistic t and find its distribution under H<sub>0</sub>
step 3: for a significance level α, make a test decision and interpret the outcome.

# Hypothesis testing

<u>Test decision</u>: given a specific significance level  $\alpha$ 

- find the critical value  $t_{1-\alpha/2}$  (two-sided test),  $t_{\alpha}$  (left-sided test) or  $t_{1-\alpha}$  (right-sided test) and compare it to the test statistic. Reject the null hypothesis if  $|t| > |t_{crit}|$ .
- calculate the (empirical) p-value and compare it to the significance level α. Reject the null hypothesis if p ≤ α.
- for two-sided test: Reject  $H_0$  if the hypothetical parameter lies outside of the associated  $1 \alpha$  confidence interval.

Interpretation: has three ingredients

The null hypothesis H<sub>0</sub> can (not) be rejected at the α significance level.

## Two ways of testing

 $\theta$  unknown parameter in the population:

1 two-sided test:

$$H_0: \theta = \bar{\theta} \qquad \qquad H_1: \ \theta \neq \bar{\theta}$$

2 one-sided test: right-sided and left-sided

$$\begin{array}{ll} H_0: \theta \leq \bar{\theta} & H_1: \ \theta > \bar{\theta} \\ H_0: \theta \geq \bar{\theta} & H_1: \ \theta < \bar{\theta} \end{array}$$

Testing problem:

The price of a cup of cappuccino in 27 cafes in Tübingen and its surroundings is investigated. The average price is given by  $\bar{y}_n = 3.10 \in$ , the standard deviation is given by  $\sigma = 0.30 \in$  and the price of a cup of cappuccino is normally distributed. We want to test whether the average price of a cappuccino  $\mu$  is significantly different from  $3 \in$ .

step 1 State the null and alternative hypothesis for the test.

$$H_0: \mu = 3$$
 and  $H_1: \mu \neq 3$ 

step 2 Set up an appropriate test statistic for the null hypothesis. What is the distribution of the test statistic under the null hypothesis?

$$Z = \frac{\bar{Y}_n - 3}{\sigma/\sqrt{n}} \underset{H_0}{\sim} \mathcal{N}(0, 1)$$

Hypothesis Testing

step 3 Calculate the test statistic and the associated *p*-value. What is the test decision on a 5% (10%) significance level?

test statistic: 
$$z = \frac{3.1-3}{0.30/\sqrt{27}} = 1.73$$

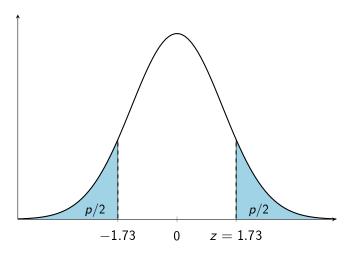
$$p = 2 P(Z > 1.73) = 2 (1 - P(Z \le 1.73)) = 2 (1 - \Phi(1.73))$$
$$= 2 (1 - 0.9582) = 2 \cdot 0.0418 = 0.0836 > \alpha = 0.05$$

Interpretation:

We cannot reject the null hypothesis  $H_0: \mu = 3$  at the 5% significance level.

But: We can reject the null hypothesis  $H_0: \mu = 3$  at the 10% significance level.

Distribution under  $H_0$ , test statistic and p-value of a two-sided test



# Two types of errors

Two types of errors:

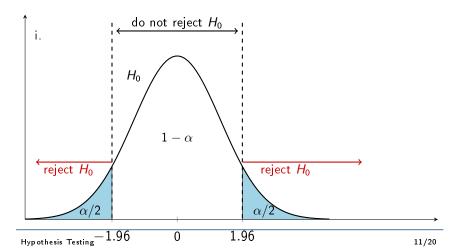
- α type | error: Probability to reject a correct null hypothesis.
- β type II error: Probability not to reject a wrong null hypothesis.

#### How "good" is the test?

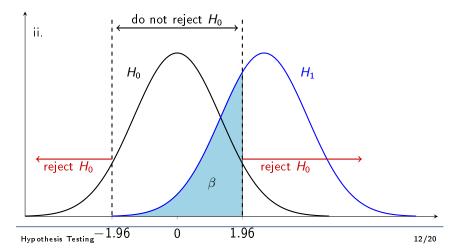
- We control the type I error, the significance level  $\alpha$ .
- The goodness of the test is measured by the **power** =  $1 \beta$ , the probability to reject a wrong null hypothesis and make a test decision.

<u>Note</u>: In practice, we do not know whether a given null hypothesis is correct or wrong in the population. We can only observe a finite sample of observations.

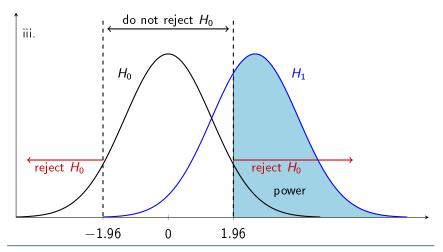
<u>Scenario</u>: Sketch the distribution of the test statistic under the null hypothesis that  $\mu = 3 \in$ . Include the rejection and non-rejection region and the type I error. Use a significance level of  $\alpha = 5\%$ .



<u>Scenario</u>: Now, assume that in reality the true price is given by  $\mu = 5 \in$ . Include a sketch of the distribution under this alternative and the type II error into the plot.



What is the power of a test? Where would we include power in the plot above?



### Two types of errors

Type I error and significance level  $\alpha$ : right-tailed test

- We select  $\alpha$ :  $P(t > c \mid H_0) = \alpha \rightarrow \text{critical value: } c = t_{1-\alpha}$ .
- We reject  $H_0$  if  $t > t_{1-\alpha}$  (with rejection area  $[t_{1-\alpha}, \infty]$ ) and don't reject  $H_0$  if  $t < t_{1-\alpha}$ .

Type II error:

• Under  $H_1$ , the most likely values of t are on the right of  $f(\bar{\theta})$ .

$$\underbrace{\beta(\theta)}_{\text{depends on } \theta, \text{ the true parameter}} = \int\limits_{-\infty}^{c} f(t,\theta) \, dt$$

ightarrow can't be controlled!

### Power of a test

Trade-off:  $\alpha$  vs  $\beta$ 

- When  $\alpha$  is selected small, the chances to reject  $H_0$  are small.  $\rightarrow$  low probability of a type I error.
- However, the probability of a type II error is large: we may fail to reject a incorrect  $H_0$ .

Power of a test:  $1 - \beta$ 

- Use the **power** of the test to assess the "goodness" of the test.
- $1 \beta$  is the probability to reject an null hypothesis given that it is incorrect in population.
- The faster the probability to reject  $H_0$  increases (steeper red line), the better when  $\theta \neq \overline{\theta}$  (two-sided test) or  $\theta > \overline{\theta}$  (right-sided test) is true in population.

# What does significant really mean?

statistical significance

- does not answer the question wether the null hypothesis is wrong or right in population
- does not indicate how (un-) likely the null hypothesis is
- only controlled by maximum probability to run into type I error  $(\alpha)$
- provides no control over probability of type II error  $(\beta)$

goal: for  $\alpha$  given

- $\rightarrow$  minimal  $\beta$
- $\rightarrow$  minimal  $\alpha + \beta$
- ightarrow maximal 1-eta

## Finite sample t-test using normality

estimated parameters  $\widehat{eta_1} \dots \widehat{eta_K}$ 

- **1** step 1: define  $H_0$ , e.g.  $H_0 : \beta_k = \overline{\beta}_k$  and  $H_1$ , e.g.  $H_1 : \beta_k \neq \overline{\beta}_k$
- 2 step 2: construct test statistic

$$t = rac{\widehat{eta_k} - ar{eta_k}}{s.e.(\widehat{eta_k})} \sim t(N-K)$$
 under  $H_0$ 

- **3** step 3: test decision choose significance level  $\alpha$
- compare test statistic t and critical value or the empirical p-value and the significance level α

### Large sample case

estimated parameters  $\widehat{eta}_1 \dots \widehat{eta}_K$ 

- **1** step 1: define  $H_0$ , e.g.  $H_0 : \beta_k = \overline{\beta}_k$  and  $H_1$ , e.g.  $H_1 : \beta_k \neq \overline{\beta}_k$
- step 2: construct test statistic and use law of large numbers (LLN) and central limit theorem (CLT)

$$\mathsf{z} = rac{\widehateta_k - areta_k}{s.e.(\widehateta_k)} \stackrel{\mathsf{a}}{\sim} \mathcal{N}(0,1)$$
 under  $H_0$ 

- **3** step 3: test decision choose significance level  $\alpha$
- compare test statistic z and critical value z<sub>crit</sub> or the empirical p-value and the significance level α

### **Confidence Interval**

Construct a  $(1 - \alpha)$  confidence interval around  $\widehat{\beta}_k$ :

In finite sample (+normality assumption), use

$$CI(\widehat{\beta}_k,\alpha) = \left[\widehat{\beta}_k - t_{1-\frac{\alpha}{2}} \cdot s.e.(\widehat{\beta}_k), \widehat{\beta}_k + t_{1-\frac{\alpha}{2}} \cdot s.e.(\widehat{\beta}_k)\right].$$

In large samples, we can also make use of the LLN and CLT and use

$$CI(\widehat{\beta}_k,\alpha) \approx \left[\widehat{\beta}_k - z_{1-\frac{\alpha}{2}} \cdot s.e.(\widehat{\beta}_k), \widehat{\beta}_k + z_{1-\frac{\alpha}{2}} \cdot s.e.(\widehat{\beta}_k)\right]$$

**Interpretation**: If  $\bar{\beta}_k$  is contained in the confidence interval,  $H_0: \beta_k = \bar{\beta}_k$  cannot be rejected on the  $\alpha$  significance level.

## LLN and CLT

For a random sample  $\{X_1, X_2, \ldots, X_n\}$  with finite  $E(X_i)$  and  $Var(X_i)$  and an appropriately large n, it holds that

Law of large numbers (LLN)

$$\lim_{n\to\infty} P\left[ \left| \frac{1}{n} \sum_{i=1}^n X_i - E(X_i) \right| > \epsilon \right] = 0 \quad \text{for any } \epsilon > 0.$$

and

Central limit theorem (CLT)

$$\sqrt{n}\left[\frac{1}{n}\sum_{i=1}^{n}X_{i}-E(X_{i})
ight]\overset{a}{\sim}\mathcal{N}(0,Var(X_{i}))$$

Hypothesis Testing